# MODIFIABLE AND NON-MODIFIABLE RISK FACTORS OF CABG PATIENTS AND PARAMETRIC ESTIMATION OF SURVIVAL PROPORTIONS OF COMPLETE POPULATIONS 

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#### Abstract

In this paper, modifiable and non-modifiable risk factors of Coronary Artery Bypass Graft Surgery (CABG) patients are considered. A new approach of complete population from its incomplete population is used. In the complete population, censored patients have been proportionally incorporated into the known died and survived patients respectively. The availability of a complete population may represent better behavior of lifetimes / survival proportions for medical investigations. Survival proportions of the CABG patients of complete populations, with respect to modifiable and non-modifiable risk factors, are obtained from suitable parametric models (Weibull and Exponential). Maximum likelihood method, in-conjunction with Davidon-Fletcher-Powell (DFP) optimization method and Cubic Interpolation method is used in estimation of survivor's proportions of the parametric models.


Keywords- CABG Patients, Complete \& Incomplete Populations, Modifiable \& Non-Modifiable Risk Factors, Parametric models (Weibull and Exponential), Maximum likelihood method, Davidon-Fletcher-Powell optimization method and Survivor's Proportions.

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## Introduction

The Coronary Artery Disease (CAD) is due to accumulation of cholesterol and other material, called plaque, within inner walls of the coronary arteries (the arteries that supply blood with oxygen and nutrients to heart muscles). As this buildup grows (in medical term, Arteriosclerosis), only less blood can flow through the arteries. Over the time heart weakens. This leads to chest pain (Angina) which is a symptom of Myocardial Infarction (MI). When the clot (thrombus) completely cuts off the hearts blood supply, this leads to permanent heart damage or heart attack (MI). Heart failure means the heart can't pump blood well to the rest of the body [10]. CAD is the leading cause of death worldwide (see William, Stephen, Van -Thomas and Robert) [36], John [18], Hansson [13], Axel, Yiwen, Dalit, Veena, Elaine, Catia, Matthew, Jonathan, Edward \& Len [3] and Sun \& Hoong [33] ). The symptoms and signs of coronary artery disease are noted in the advanced stage of disease, most individuals with CAD show no evidence of disease for decades as the disease progresses before the first
onset of symptoms, often there is a sudden heart. The disease is the most common cause of sudden death of men and women over 20 years of age [30]. The medical scientists; William, Ellis, Josef, Ralph and Robert [36], Heymann [15], Goldstein, Adams, Alberts, Appel, Brass, Bushnell, Culebras, DeGraba, Gorelick \& Guyton [11] and Jennifer [17] are of the opinion that CABG is an effective treatment option for CAD patients. The medical research organizations like Heart and Stroke Foundation Canada [14] and American Heart Association [2] have classified risk factors of CABG patients as modifiable and non-modifiable. Modifiable risk factors are those factors, which may be controlled by changing life style or taking medications to reduce cardiovascular risk. These risk factors include high blood pressure (hypertension), diabetes, smoking, high cholesterol, sedentary lifestyle and obesity. Nonmodifiable risk factors are those factors, which may not be changed. These factors include age, gender and family history (Genetic predisposition). William, Ellis, Josef, Ralph and Robert [35] in 1995 carried out the survival study on incomplete popula-
tion (progressive censoring of type 1) of CABG patients comprising 2011 patients using Kaplan Meier method [19]. The patients were grouped with respect to Male, Female, Age, Hypertension, Diabetes, and Ejection Fraction, Vessels, Congestive Heart Failure, Elective and Emergency Surgery. The patients were undergone through a first re-operation at Emory University hospitals from 1975 to 1993. This study also comprises the same data set of 2011 patients. The details of patients are given in the article [35].
In this paper we present survival study of the CABG patients with respect to some modifiable risk factors, Hypertension \& Diabetes and non-modifiable risk factors. Gender and Age. Khan, Saleem and Mahmud in the article [21] proposed a procedure, to make an incomplete population (IP) a complete population (CP). The differences between the means of survival proportions of the CABG patients, obtained by using different survival models (Weibull, Expontial etc) are statistically at $5 \%$ level of significance for details see the article [21].
The importance of parametric models for analysis of lifetime date has been indicated by Mann, Schefer and Singpurwala [28], Nelson [29], Cyrus [8], Lawless [25], Klein \& Moeschberger [22] and Sridhar and Mun Choon Chan [32]. The Exponential distribution model has been used by Lee, Kim and Jung [27] in medical research for survival data of patients. The Weibull distribution model has been used for survival analysis by Cohen [6], Gross and Clark [12], Bunday [5], Crow [7], Klein \& Moeschberger [22], Lawrencce [26], Abrenthy [1], Hisada \& Arizino [16], Lawless [25], David \& Mitchel [9] and Lang [23]. In particular, the survival study of chronic diseases, such as AIDS and Cancer, has been carried out by Bain and Englehardt [4], Khan \& Mahmud [20], Klein \& Moeschberger [22], Lawless [25] and Swaminathan and Brenner [34] using Exponential and Weibull distributions. Lanju \& William [24] used Weibull distribution to human survival data of patients with plasma cell and in response-adaptive randomization for survival trials respectively. Lee, Kim and Jung [27] used exponential in medical research for survival data of the patients. Khan, Saleem and Mohmud [21] concluded that the survival data of the CABG patients has been best modeled by the Weibull and Exponential distributions. In this paper, the survivor proportions of the CABG patients are obtained for complete population of the CABG patients by parametric models (Weibull and Exponential), using data of CABG patients groups: Gender, Age, Hypertension and Diabetes. Maximum likelihood method, in-conjunction with DFP optimization method and Cubic Interpolation method is used. A subroutine for maximizing log-likelihood function of each model is developed in FORTRAN program to obtain the estimates of the parameters of the model. The survival proportions of $C P$ of the CABG patients with respect to modifiable and non-modifiable risk factors risk factors are presented, discussed and concluded.

## Methodology

Khan, Saleem and Mahmud [21] mentioned that the method proposed by Kaplan Meier [19] and latter discussed by William [35] in
1995 and Lawless [25] is: $S(t)=\prod_{j: t_{j}<t}\left(1-\frac{d_{j}}{n_{j}}\right)$,where $d_{j}$
and ${ }^{n}{ }_{j}$ are the number of items (individuals / patients) failed
(died individuals) and number of individuals at risk at time $t_{j}$ that is, the number of individuals survived and uncensored at time
$t_{j-1}$. This method does not take into account the censored individuals ${ }^{c_{j}}$ completely and thus the analysis is performed on (IP) incomplete population Further, Khan, Saleem and Mahmud [21] proposed that the censored individuals ${ }^{c_{j}}$ could be taken into account. The inclusion of splitted-censored individuals, $c_{j}$ proportionally $\left[\left(1-\frac{d_{j}}{n_{j-1}-c_{j}}\right) \times c_{j}\right.$ and $\left.\left(\frac{d_{j}}{n_{j-1}-c_{j}}\right) \times c_{j}\right]$ into known survived, $n_{j}$ and died individual's $d_{j}$ respectively makes the population complete. Thus the survival analysis may be
performed on the complete population
(CP) and Khan [31] also mentioned the form of likelihood function proposed by Klein \& Moeschberger [22] and Lawless [25], for a survival model, in the presence of censored data. The maximum likelihood method works by developing a likelihood function based on the available data and finding the estimates of parameters of a probability distribution that maximizes the likelihood function. This may be achieved by using iterative method: see Bunday \& AlMutwali [5] and Khan \& Mahmud [20] The likelihood function for all observed died and censored individuals is of the form:

$$
L(t ; \underline{\theta})=\prod_{i=1}^{n}\left[f\left(t_{i} ; \underline{\theta}\right)\right]^{f_{t_{i}}} \prod_{i=1}^{n}\left[S\left(t_{i} ; \underline{\theta}\right)\right]^{c_{t_{i}}}, \text { where } f_{t_{i}} \& c_{t_{i}}
$$

are the number of died \& censored individuals in interval $i$ each of length $t, f(t ; \underline{\theta})$ is pdf in a parametric model with survivor function, $S(t ; \underline{\theta})$ \& hazard function, $h(t ; \underline{\theta})$ and $\underline{\theta}$ is vector of parameters say $\underline{\theta}=(\alpha, \beta)$ of the model. To obtain maximum likelihood estimates of parameters of a parametric model using DFP optimization method, we take negative log on both the sides of above equation and therefore by setting $l=-\ln (L(t ; \underline{\theta}))$, we get:

$$
\begin{align*}
l & =-\sum_{i=1}^{n} f_{t_{i}} \ln f\left(t_{i} ; \underline{\theta}\right)-\sum_{i=1}^{n} c_{t_{i}} \ln \left(S\left(t_{i} ; \underline{\theta}\right)\right) \\
& =-\sum_{i=1}^{n} f_{t_{i}} \ln h\left(t_{i} ; \underline{\theta}\right)-\sum_{i=1}^{n}\left(f_{t_{i}}+c_{t_{i}}\right) \ln \left(S\left(t_{i} ; \underline{\theta}\right)\right)
\end{align*}
$$

$$
f\left(t_{i} ; \underline{\theta}\right)=h\left(t_{i} ; \underline{\theta}\right) S\left(t_{i} ; \underline{\theta}\right)
$$

Where, the first sum is for failure and the second sum is for all censored individuals.
Setting $N_{t_{i}}=\left(f_{t_{i}}+c_{t_{i}}\right)$, where $N_{t_{i}}$ represents total no of individuals at time ${ }^{t_{i}}$ we get:

$$
\begin{equation*}
l=-\sum_{i=1}^{n} f_{t_{i}} \ln h\left(t_{i} ; \underline{\theta}\right)-\sum_{i=1}^{n}\left(N_{t_{i}}\right) \ln \left(S\left(t_{i} ; \underline{\theta}\right)\right) \tag{1}
\end{equation*}
$$

In this study time is partitioned into intervals, which are of unit length $t$ starting from zero. Moreover, failures and censoring of the patients occur in each interval $i$ of equal length of time $t$,

$$
i=1,2, \ldots, 12
$$

For complete population the term for censored observations is dropped from the likelihood function.

## 3. Application

Khan, Saleem and Mahmud [21] presented detail application of above methodology for parametric model (Weibull distribution). Same procedure is followed for second parametric model (Exponential distribution) considered in this article. The methodology is reproduced here. The probability density function (pdf) of
Weibull distribution is: $f(t ; \underline{\theta})=\left(\frac{\beta}{\alpha}\right)\left(\frac{t}{\alpha}\right)^{\beta-1} e^{-\left(\frac{t}{\alpha}\right)^{\beta}}$, where $\underline{\theta}$
is vector of parameters $\alpha$ and $\beta ; \alpha$ is a scale parameter and $\beta$ is a shape parameter; $\alpha, \beta$ and $t>0$.
The survival and hazard functions of Weibull distribution are:

$$
S(t ; \underline{\theta})=e^{-\left(\frac{t}{\alpha}\right)^{\beta}} \text { and } h(t ; \underline{\theta})=\left(\frac{\beta}{\alpha}\right)\left(\frac{t}{\alpha}\right)^{\beta-1}
$$

For incomplete population replacing values of the survival and hazard functions of Weibull distribution in equation (1), we get

$$
\begin{align*}
& l=-\sum_{i=1}^{n} f_{t_{i}} \ln \left[\left(\frac{\beta}{\alpha}\right)\left(\frac{t_{i}}{\alpha}\right)^{\beta-1}\right]-\sum_{i=1}^{n} N_{t_{i}} \ln \left(e^{-\left(\frac{t_{i}}{\alpha}\right)^{\beta}}\right) \\
& =-(F) \ln \left(\frac{\beta}{\alpha}\right)-(\beta-1) \sum_{i=1}^{n} f_{t_{i}} \ln \left(\frac{t_{i}}{\alpha}\right)+\sum_{i=1}^{n} N_{t_{i}}\left(\frac{t_{i}}{\alpha}\right)^{\beta} \text { or }  \tag{2}\\
& \qquad F=\sum_{i=1}^{n} f_{t_{i}} \quad \text { is the total number of failures in a given time. }
\end{align*}
$$ Differentiating (2) with respect to $\alpha$ and $\beta$ and simplifying we get

$$
\begin{align*}
& \frac{\partial l}{\partial \alpha}=F\left(\frac{\beta}{\alpha}\right)-\left(\frac{\beta}{\alpha}\right) \sum_{i=1}^{n} N_{t_{i}}\left(\frac{t_{i}}{\alpha}\right)^{\beta}  \tag{3}\\
& \frac{\partial l}{\partial \beta}=-\frac{F}{\beta}-\sum_{i=1}^{n} f_{t_{i}} \ln \left(\frac{t_{i}}{\alpha}\right)+\sum_{i=1}^{n} N_{t_{i}}\left(\frac{t_{i}}{\alpha}\right)^{\beta} \ln \left(\frac{t_{i}}{\alpha}\right) \tag{4}
\end{align*}
$$

By using (2), (3) and (4) in the DFP optimization method, we find the parameters estimates for which value of the likelihood function is maximum. For complete population we drop the term for cen-
sored observations from likelihood function. Same procedure is followed for Exponential model. FORTRAN program for the parameters estimation of both the models is developed. The optimal estimates of the scale and shape / location parameters ( $\alpha$ and $\beta$ respectively) of Weibull and Exponential distributions distribution using CP of groups: Hypertension (Absent \& Present), Diabetes (Absent \& Present), Male \& Female and Age groups (Less than 50 years (I), 50 to 59 years (II), 60 to 59 years (III) and 70 \& above years (IV)) are obtained by maximizing the log-likelihood function. The $t$-ratios of the parameters are given in parenthesis. The values of parameters estimates, $t$-ratios, log-likelihood function and variance-covariance matrix are given below: -

## Weibull Distribution

a] Non-Modifiable factors of CABG Patients (Gender and Age)

## Male \&Female and Age Groups I, II, III \& IV (Male and Female CABG Patients)

The survival proportions of male \& female and age groups I, II, III \& IV of CABG patients are obtained using Weibull distribution (two parametric; $\alpha$ and $\beta$ ) as explained earlier. The optimal estimates of the parameters obtained by maximizing the log-likelihood function are given below in table 1.
The estimated values of scale parameter $\alpha>0$ and shape parameter $\beta>0$ for male \& female and age groups I, II, III \& IV of CABG patients are given in the table 1 along with $t$-ratios in the parenthesis, indicating that the estimates of scale and shape parameters are significant at $5 \%$ level of significance. The estimated value of $\beta$ is greater than 1 for the male \& female and age groups. I, II, III \& IV, of CABG patients, indicate the increasing failure rate with time. The positive or negative values of co-variances indicates that the movements of $\hat{\alpha}$ and $\hat{\beta}$ are in the same or opposite directions respectively.
The estimated survival proportions of male \& female and age groups I, II, III \& IV of CABG patients and respective graphs are given in table 2.

## Discussion

The graph of survival proportions obtained by using Weibull distribution for male and female groups of CABG patients indicates that for the difference between the survival proportions is small at the start, continuously but slowly increasing, whereas the survival proportions of female group of the CABG patients are lower than those of male group of the CABG patients. The graph of survival proportions obtained by using Weibull distribution for age groups I, II, III and IV of CABG patients indicates that for the age groups I and II, the difference between survival proportions of is small, whereas for the age groups III and IV the difference between survival proportions of is small at the start, and increasing continuously. The survival proportions of the age group IV are comparatively lowest in four age groups.

## Conclusion

The survival proportions of female group of the CABG patients are lower than those of female group of the CABG patients. The survival proportions of the age group IV are comparatively lowest in
four age groups, which indicates increasing failure rate with the increase in age.

## b] Modifiable factors of CABG Patients (Hypertension and Diabetes

Hypertension Absent ( ${ }^{\hat{y}_{t}}$ ) HYa \& Present $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{HYp}}$ and Diabetes Absent ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{Da}} \& \operatorname{Present}\left({ }^{\hat{y}_{t}}\right)_{\mathrm{Dp}}$
The survival proportions ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYa}}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYp}}$ and ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{Da}}$ \& ( ${ }^{\hat{y}_{t}}$ ) Dp groups of the CABG patients are obtained using Weibull distribution as explained earlier. The optimal estimates of parameters obtained by maximizing the log-likelihood function are given in table 3.
The estimated values of scale parameter $\alpha>0$ and shape parameter $\beta>0$ for $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients are given in the table 3 along with $t$-ratios in the parenthesis, indicating that the estimates of scale and shape parameters are significant at $5 \%$ level of significance. The estimated value of $\beta$ is greater than 1 for $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients, which indicates the increasing failure rate with time. The positive or negative values of co-variances indicates that the movements
of $\hat{\alpha}$ and $\hat{\beta}$ are in the same or opposite directions respectively.

The estimated survival proportions of $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients and respective graphs are given in table 4.

## Discussion

The graph of survival proportions (obtained by using Weibull distribution) for $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients shows that the difference between the survival proportions or $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ is small at the start and continuously increasing.

## Conclusion

The survival proportions (obtained by using Weibull distribution) of the CABG patients in the presence of hypertension and diabetes are lower than in the absence of hypertension and diabetes.
Finally, as a whole the survival proportions (obtained by using Weibull distribution), of the CABG patients groups ( $H Y_{a} \& H Y_{p}$ and $\left.D_{a} \& D_{p}\right)$ in the presence of diabetes are lowest.

## Exponential Distribution

## a] Non-Modifiable factors of CABG Patients (Gender and

 Age)Male \&Female and Age Groups (Male and Female CABG Patients)
Survival Proportions $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{M}}$ \& $\left({ }^{\hat{y}_{t}}\right) \mathrm{F}$ of Male \& Female and ( $\left.\left.\hat{y}_{t}\right)_{I,( }{ }^{\hat{y}_{t}}\right)_{I I I}\left({ }^{\hat{y}_{t}}\right)_{I I} \&\left({ }^{\hat{y}_{t}}\right)_{\text {IV }}$ of Age Groups I, II, III \&IV respectively of the CABG Patients from exponential distribution (as parametric,a ; keeping $\beta=1$ ) as explained earlier. The optimal esti-
mates of the parameters obtained by maximizing the log-likelihood function are given in table 5 .
The estimated values of scale parameter $\alpha>0$ for male \& female and age groups I, II, III IV of CABG patients are given in the table 5 along with $t$-ratios in the parenthesis, indicating that the estimates of scale are significant at $5 \%$ level of significance.
The estimated survival proportions of male \& female and age groups I, II, III \& IV of CABG patients and respective graphs are given in table 6.

## Discussion

The graph of survival proportions obtained by using exponential distribution for male and female groups of CABG patients indicates that the difference between the survival proportions is small at the start and increasing continuously, whereas the survival proportions of female group of the CABG patients are lower than those of male group of the CABG patients. The graph of survival proportions obtained by using exponential distribution for age groups I, II, III and IV of CABG patients indicates that for the age groups I and II, the difference between survival proportions of is small, whereas for the age groups III and IV the difference between survival proportions of is small at the start, and increasing continuously. The survival proportions of the age group IV are comparatively lowest in four age groups.

## Conclusion

The survival proportions of female group of the CABG patients are lower than those of female group of the CABG patients. The survival proportions of the age group IV are comparatively lowest in four age groups, which indicates increasing failure rate with the increase in age.

## b] Modifiable factors of CABG Patients (Hypertension and Diabetes)

The survival proportions ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYa} \&}$ ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYp}}$ and ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{Da}}$ \& $\left({ }^{\hat{y}_{t}}\right)$ $D_{p}$ of the $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients are obtained using Exponential distribution as explained earlier.
The optimal estimates of parameters obtained by maximizing the log-likelihood function are given below in table 6.
The estimated values of scale parameter $\alpha>0$ for $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients are given in the table 7 along with $t$-ratios in the parenthesis, indicating that the estimates of scale and shape parameters are significant at $5 \%$ level of significance.
The estimated survival proportions of $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients and respective graphs are given in table 8.

## Discussion

The graph of survival proportions (obtained by using exponential distribution) for $H Y_{a}$ \& $H Y_{p}$ and $D_{a} \& D_{p}$ groups of CABG patients shows that the difference between the survival proportions or $\mathrm{HY}_{\mathrm{a}}$ \& $H Y_{p}$ and $D_{a} \& D_{p}$ is small at the start and increasing continuously.

## Conclusion

The survival proportions (obtained by using Exponential distribu-
tion) of the CABG patients in the presence of Hypertension and Diabetes are lower than in the absence of Hypertension and

## diabetes.

Finally, as a whole the survival proportions (obtained by using Exponential distribution), of the CABG patients groups $\left(H Y_{a}\right.$ \& $H Y_{p}$ and $\left.D_{a} \& D_{p}\right)$ in the presence of Diabetes are lowest.

## References

[1] Abernathy R.B. (1998) The New Weibull Handbook. 3rd ed. SAE Publications, Warrendale.PA.
[2] American Heart Association Dallas, Texas (2007) Heart Disease and Stroke Statistics.
[3] Axel Vise, Yiwen Zhu, Dalit May, Veena Afza, Elaine Gong, Catia Attanasio, Matthew J. Blow, Jonathan C. Cohen, Edward M. Rubin \& Len A. Pennacchio (2010) Targeted deletion of the 9p21 non-coding coronary artery disease risk interval in mice.
[4] Bain L.J. and Englehardt M. (1991) Statistical Analysis of Reliability and Life-Testing Models: Theory and Methods, 2nd ed.
[5] Bunday B.D. and Al-Mutwali I.A. (1981) IEEE Trans, Reliability, R-30(4), 367-339.
[6] Cohen A.C. (1965) Technometrics., 7(4), 579-588.
[7] Crow L.H. (1982) Technometrics, 24(1), 67-72.
[8] Cyrus R. Mehta (1981) Biometrika 68(3), 669-675.
[9] David G. Kleinbaum, Mitchel Klein (2005) Survival analysis: a self-learning text.
[10]Dorlands Medical Dictionary (2009) Coronary Artery Disease.
[11]Goldstein L., Adams R., Alberts M., Appel L., Brass L., Bushnell C., Culebras A., DeGraba T., Gorelick P., Guyton J. (2006) American Journal of Ophthalmology: American Heart Association., 142(4), 716-716.
[12]Gross A.J. and Clark V.A. (1975) Survival Distribution: Reliability Applications in the Biomedical Sciences Wiley.
[13]Hansson Göran K. (2005) Inflammation, Atherosclerosis, and Coronary Artery Disease. 352(16), 1685-1695.
[14]Heart and Stroke Foundation Canada (1997) Heart Disease and Stroke Statistics.
[15]Heymann C. Von Heymann (2002) Journal of Cardiothoracic and Vascular Anesthesia, 16(5), 615-616.
[16]Hisada \& Arizino (2002) IEEE Transactions, 51(3), 331-336.
[17]Jennifer Heisler R.N. (2008) After Coronary Artery Bypass Graft Surgery-Recovering From Open Heart Surgery.
[18]John H. Lemmer (2003) Hand Book of Patient Care in Cardiology Surgery, Lippincott Williams \& Wilkins.
[19]Kaplan E.L., Meier P. (1958) Nonparametric estimations from incomplete observations.
[20]Khan K.H. and Mahmud Z. (1999) J. American Assoc., 53, 457-481.
[21]Khan K.H., Saleem M. and Mahmud Z. (2011) Survival Proportions of CABG Patients: A New Approch, 3(3).
[22]Klein P.J. and Moeschberger L.M. $(1997,2003)$ Survival Analysis Techniques for Censored and Truncated Data.
[23]Lang Wu (2010) Mixed effects models for complex data.
[24]Lanju Zhang and William F. Rosenberger (2007) Journal of the Royal Statistical Society: Series C, 56(2), 153-165.
[25]Lawless Jerald F. Lawless $(1982,2003)$ Statistical Models and Methods for Lifetime Data, John Wiley and Sons.
[26]Lawrence M. Leemis (1995) Reliability Probabilistic Modela and Statistical Methods.
[27]Lee Jaeyong, Kim Jinseog and Jung Sin-Ho (2006) Lifetime Data Analysis, 13(1).
[28]Maan N.R., Schafer R.E. and Singpurawalla N.D. (1974) Method for Statistical Analysis of reliability and Lifetime Data. Wiley.
[29]Nelson W. (1982) Applied Life Data Analysis.
[30]Rao Venkata and Kiran Ravi (2011) J. Cardiovasc. Dis. Res., 2(1), 57-60.
[31]Saleem M., Mahmud Z. and Khan K.H. (2012) American Journal of Statistics and Mathematics.
[32]Sridhar K.N., Mun Choon Chan (2009) Modeling link lifetime data with parametric regression models in MANETs. IEEE 13 (12).
[33]Sun Zhonghua and Hong. Ng. Kwan- (2011) World J. Cardiol., 3(9), 303-310.
[34]Swaminathan R. and Brenner H. $(1998,2011)$ Statistical methods for Cancer Survival Analysis, 1 \& 2.
[35]William S. Weintraub, Ellis L. Jonees, Josef M. Craver, Ralph Grossedwald, Robert A. Guyton $(1995,1997)$ American Heart Association.
[36]William S. Weintraub, Stephen D. Clements, Van-Thomas Crisco L., Robert A. Guyton (2003) American Heart Association.

Table 1- Estimates of Parameters of Weibull Distribution Using Data of Male \& Female and Age Groups (I, II, III \& IV) of CABG Patients

| Parameters | Male Group |  | Female Group |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Estimates/ (tratio) | Gradients | Estimates/ (tratio) | Gradients |
| $\alpha$ | $\begin{aligned} & 16.2415 \\ & (30.64) \end{aligned}$ | $-7.24 \times 10^{-11}$ | $\begin{aligned} & 13.3967 \\ & (16.03) \end{aligned}$ | $2.67 \times 10^{-9}$ |
| $\beta$ | $\begin{aligned} & 1.1598 \\ & (31.68) \end{aligned}$ | $-2.71 \times 10^{-10}$ | $\begin{aligned} & 1.1689 \\ & (15.41) \end{aligned}$ | $-4.47 \times 10^{-8}$ |
| Log-Likelihood | 3302.079 |  | 728.6315 |  |
| VarianceCovariance | ${ }^{0.281}$ | $-8.4 \times 10^{-3}$ | $(0.698$ | $\left.-1.62 \times 10^{-2}\right)$ |
| Matrix | $-8.4 \times 10^{-3}$ | $1.34 \times 10^{-3}$ ) | $-1.62 \times 10^{-2}$ | $5.75 \times 10^{-3}$ |


| Parameters | Age Groups |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimates/ (tratio) | Gradients | Estimates/ (tratio) | Gradients | Estimates/ (t-ratio) | Gradients | Estimates/ (t-ratio) | Gradients |
| $\alpha$ | $\begin{aligned} & 19.918 \\ & (10.31) \end{aligned}$ | $1.5 \times 10^{-9}$ | $\begin{aligned} & 20.966 \\ & (15.25) \end{aligned}$ | $5.65 \times 10^{-11}$ | $\begin{aligned} & 11.547 \\ & (29.81) \end{aligned}$ | $2.61 \times 10^{-10}$ | $\begin{aligned} & 8.758 \\ & (25.39) \end{aligned}$ | $-1.96 \times 10^{-8}$ |
| $\beta$ | $\begin{aligned} & 1.253 \\ & (10.63) \end{aligned}$ | $4.43 \times 10^{-8}$ | $\begin{aligned} & 1.197 \\ & (16.76) \end{aligned}$ | $-4.17 \times 10^{-8}$ | $\begin{aligned} & 1.330 \\ & (25.00) \end{aligned}$ | $9.74 \times 10^{-9}$ | $\begin{aligned} & 1.414 \\ & (20.47) \end{aligned}$ | $1.97 \times 10^{-8}$ |
| Log-Likelihood | 415.9276 |  | 1043.354 |  | 1783.982 |  | 997.2853 |  |
| VarianceCovariance Matrix | $\left(\begin{array}{cc}3.735 & -0 \\ -0.131 & 1\end{array}\right.$ | $\left.\times 10^{-2}\right)$ | $\left(\begin{array}{l}1.889 \\ -5.83 \times 10^{-2}\end{array}\right.$ | $\left.\begin{array}{c}-5.83 \times 10^{-2} \\ 5.1 \times 10^{-3}\end{array}\right)$ | $\left(\begin{array}{l}0.150 \\ -3.56 \times 10^{-3}\end{array}\right.$ | $\left.\begin{array}{l}-3.56 \times 10^{-3} \\ 2.83 \times 10^{-3}\end{array}\right)$ | $\left(\begin{array}{l}0.119 \\ 2.73 \times 10^{-3}\end{array}\right.$ | $\left.\begin{array}{l}2.73 \times 10^{-3} \\ 4.77 \times 10^{-3}\end{array}\right)$ |

Table 2- Survival Proportions ( $\left.{ }^{\hat{y}_{t}}\right)_{M}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{F}$ of Male \& Female and $\left.\left.\left({ }^{\hat{y}_{t}}\right)_{I,( } \hat{y}_{t}\right)_{I I,( }{ }^{\hat{y}_{t}}\right)_{I I I}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{I V, \text { of Age Groups I, II, III \& IV re- }}$ spectively CABG Patients from Weibull Distribution and Respective Graphs are as under.

| Years (t) | Male \& Female Groups |  | Age Groups |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | I | II |  |  |
|  | $\hat{y}_{t}$ | $\hat{y}_{t}$ | $\hat{y}_{t}$ | $\hat{y}_{t}$ | $\hat{y}_{t}$ | $\hat{y}_{t}$ |
| 0 | 1 | 1 | 1 | 1 | 1 | 1 |
| 1 | 0.961 | 0.953 | 0.976 | 0.974 | 0.962 | 0.954 |
| 2 | 0.915 | 0.897 | 0.945 | 0.941 | 0.907 | 0.883 |
| 3 | 0.868 | 0.840 | 0.911 | 0.907 | 0.846 | 0.802 |
| 4 | 0.821 | 0.783 | 0.874 | 0.871 | 0.783 | 0.718 |
| 5 | 0.774 | 0.729 | 0.837 | 0.835 | 0.720 | 0.636 |
| 6 | 0.729 | 0.676 | 0.800 | 0.799 | 0.658 | 0.556 |
| 7 | 0.686 | 0.626 | 0.763 | 0.764 | 0.598 | 0.482 |
| 8 | 0.644 | 0.578 | 0.727 | 0.729 | 0.541 | 0.414 |
| 9 | 0.604 | 0.533 | 0.691 | 0.695 | 0.487 | 0.353 |
| 10 | 0.565 | 0.491 | 0.656 | 0.662 | 0.437 | 0.299 |
| 11 | 0.529 | 0.451 | 0.621 | 0.630 | 0.391 | 0.251 |
| 12 | 0.494 | 0.415 | 0.588 | 0.598 | 0.349 | 0.209 |



Table 3- Estimates of Parameters of Weibull Distribution Using Data of $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ Groups of CABG patients.

| Parameters | Age Groups |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\xrightarrow[\text { Estimates! }]{H Y_{a}}$ |  | HYp |  | $D_{a}$ |  | $D_{p}$ |  |
|  | $\begin{gathered} \text { Estimates/ } \\ \text { (t-ratio) } \end{gathered}$ | Gradients | $\begin{aligned} & \text { Estimates/ } \\ & \text { (tratio) } \end{aligned}$ | Gradients | $\begin{aligned} & \text { Estimates/ } \\ & \text { (tratio) } \end{aligned}$ | Gradients | Estimates/ (t-ratio) | Gradients |
| $\alpha$ | 19.476 | $2.48 \times 10-10$ | 11.659 | $-1.21 \times 10^{-9}$ | 18.149 | $-2.48 \times 10-10$ | 9.736 | $3.28 \times 10^{-10}$ |
| $\alpha$ | (19.92) | $2.48 \times 10$ | (31.73) | $-1.21 \times 10^{-9}$ | (26.01) | -2.48 $\times 10^{-10}$ | (24.57) | $3.28 \times 10$ |
| $\beta$ | 1.142 | -6.22 $\times 10^{-10}$ | 1.316 | $-7.1 \times 10^{-8}$ | 1.154 | -1.36 x 10-9 | 1.380 | -3.91 x 10-8 |
| $\beta$ | (22.14) | $-6.22 \times 10$ | (26.81) | -7.1×108 | (27.83) | -1.36 x 10 | (20.07) | -3.91 $\times 10$ |
| Log-Likelihood | 1743.991 |  | 2054.818 |  | 2688.629 |  | 1023.245 |  |
| Variance- | (0.956 | $-2.74 \times 10^{-2}$ | (0.135 | $-2.93 \times 10^{-3}$ |  | -1.45*10 ${ }^{-2}$ ) | 0.157 | $1.01 \times 10^{-4}$ ) |
| Covariance Matrix | ${ }_{-2.74 \times 10^{-2}}$ | $2.66 \times 10^{-3}$ ) | .$^{-2.93 \times 10^{-3}}$ | $2.41 \times 10^{-3}$ ) | -1.45*10 ${ }^{-2}$ | $1.72 \times 10^{-3}$ ) | $1.01 \times 10^{-4}$ | $4.72 \times 10^{-3}$ ) |

Table 4- Survival proportions ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYa}}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{HYp}}$ and ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{Da}}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{\mathrm{Dp}}$ of the Groups of CABG Patients from Weibull Distribution and Respective Graphs

|  | $H Y_{a}$ | $H Y_{p}$ | $D_{a}$ | $D_{p}$ |
| :---: | :---: | :---: | :---: | :---: |
| Years (t) | $\left(^{\hat{y}_{t}}\right)_{\mathrm{HYa}}$ | $\left(^{\hat{y}_{t}}\right)_{H Y_{P}}$ | $\left({ }^{\hat{y}_{t}}\right)_{\text {Da }}$ | $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{DP}}$ |
| 0 | 1 | 1 | 1 | 1 |
| 1 | 0.966 | 0.961 | 0.965 | 0.957 |
| 2 | 0.928 | 0.906 | 0.924 | 0.893 |
| 3 | 0.888 | 0.845 | 0.882 | 0.821 |
| 4 | 0.848 | 0.783 | 0.839 | 0.746 |
| 5 | 0.809 | 0.720 | 0.797 | 0.671 |
| 6 | 0.770 | 0.659 | 0.756 | 0.598 |
| 7 | 0.732 | 0.600 | 0.716 | 0.530 |
| 8 | 0.696 | 0.543 | 0.678 | 0.466 |
| 9 | 0.660 | 0.491 | 0.640 | 0.407 |
| 10 | 0.626 | 0.441 | 0.605 | 0.354 |
| 11 | 0.594 | 0.396 | 0.570 | 0.306 |
| 12 | 0.562 | 0.354 | 0.537 | 0.263 |



Diabetes Absent \& Present Group


Table 5- Estimates of Parameters of Exponential Distribution Using Data of Male \& Female and Age Groups (I, II, III \& IV) of CABG Patients

| Parameter | Male Group <br> Estimates/ <br> (tratio) | Gradients | Female Group <br> Estimates/ <br> (t-ratio) | Gradients |
| :--- | :--- | :--- | :--- | :--- |
|  | $5.72 \times 10^{-2}$ <br> $(7.1)$ | $4.07 \times 10^{-7}$ | $7.1 \times 10^{-2}$ <br> $(3.76)$ | $-1.0 \times 10^{-9}$ |
| $\alpha$ | 3312.51 |  | 731.3791 |  |
| Log-Likelihood | $6.49 \times 10^{-5}$ |  | $3.56 \times 10^{-4}$ |  |
| Variance |  |  |  |  |


| Parameters | Age Groups |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 |  | II |  | III |  | IV |  |
|  | Estimates/ (tratio) | Gradients | $\begin{aligned} & \text { Estimates/ } \\ & \text { (t-ratio) } \end{aligned}$ | Gradients | $\begin{gathered} \text { Estimates/ } \\ \text { (tratio) } \end{gathered}$ | Gradients | $\begin{aligned} & \text { Estimates/ } \\ & \text { (tratio) } \end{aligned}$ | Gradients |
| $\alpha$ | $\begin{gathered} 4.31 \times 10^{-2} \\ (2.37) \end{gathered}$ | $2.11 \times 10^{-7}$ | $\begin{gathered} 4.18 \times 10^{-2} \\ (3.53) \end{gathered}$ | $1.60 \times 10^{-7}$ | $\begin{gathered} 8.19 \times 10^{-2} \\ (6.3) \end{gathered}$ | $5.07 \times 10^{-7}$ | $\begin{gathered} 0.11688 \\ (5.69) \end{gathered}$ | $-1.59 \times 10^{-7}$ |
| Log-Likelihood | 418.5939 |  | 1047.648 |  | 1807.052 |  | 1019.492 |  |
| Variance | $3.31 \times 10^{-4}$ |  | $1.40 \times 10^{-4}$ |  | $1.69 \times 10^{-4}$ |  | $4.22 \times 10^{-4}$ |  |

 spectively of the CABG Patients from Exponential Distribution and Respective Graphs are as under.



Table 7- Estimates of Parameters of Exponential Distribution Using Data of $H Y_{a} \& H Y_{p}$ and $D_{a} \& D_{p}$ Groups of CABG patients.

| Parameter |  |  |  |  |  |  |  | Gradients |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimates/ (tratio) | Gradients | Estimates/ (t-ratio) | Gradients | Estimates/ (tratio) | Gradients | Estimates/ (t-ratio) |  |
| $\alpha$ | $\begin{aligned} & 4.71 \times 10^{-2} \\ & (6.31) \end{aligned}$ | $2.09 \times 10^{-7}$ | $\begin{aligned} & 8.15 \times 10^{-2} \\ & (5.37) \end{aligned}$ | $-8.81 \times 10^{-8}$ | $\begin{aligned} & 5.07 \times 10^{-2} \\ & (4.89) \end{aligned}$ | $1.36 \times 10^{-7}$ | $\begin{aligned} & 0.1016 \\ & (6.74) \end{aligned}$ | -9.42 $\times 10^{-7}$ |
| Log-Likelihood | 1748.113 |  | 1465.947 |  | 2696.202 |  | 1041.88 |  |
| Variance | $9.26 \times 10^{-5}$ |  | $1.46 \times 10^{-4}$ |  | $6.45 \times 10^{-5}$ |  | $3.58 \times 10^{-4}$ |  |

Table 8-Survival proportions ( $\left.{ }^{\hat{y}_{t}}\right)_{H Y a}$ \& $\left(^{\hat{y}_{t}}\right)_{H Y_{p}}$ and $\left({ }^{\hat{y}_{t}}\right)_{D a}$ \& ( $\left.{ }^{\hat{y}_{t}}\right)_{D p}$ of Groups $H Y_{a} \& H Y_{p}$ and $D_{a}$ \& $D_{p}$ of CABG patients from Exponential Distribution and Respective Graphs.

| Years (t) | HYa | HYp | $D_{a}$ | $D_{p}$ |
| :---: | :---: | :---: | :---: | :---: |
| 0 | $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{HYa}}$ | $\left({ }^{\hat{y}_{t}}\right)_{\text {HYp }}$ | $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{Da}}$ | $\left({ }^{\hat{y}_{t}}\right)_{\mathrm{Dp}}$ |
| 1 | 0.954 | 0.921 | 0.950 | 0.903 |
| 2 | 0.910 | 0.849 | 0.903 | 0.816 |
| 3 | 0.868 | 0.783 | 0.859 | 0.737 |
| 4 | 0.828 | 0.721 | 0.816 | 0.666 |
| 5 | 0.790 | 0.665 | 0.776 | 0.601 |
| 6 | 0.753 | 0.613 | 0.737 | 0.543 |
| 7 | 0.719 | 0.565 | 0.701 | 0.491 |
| 8 | 0.686 | 0.520 | 0.666 | 0.443 |
| 9 | 0.654 | 0.480 | 0.633 | 0.400 |
| 10 | 0.624 | 0.442 | 0.602 | 0.362 |
| 11 | 0.595 | 0.407 | 0.572 | 0.327 |
| 12 | 0.568 | 0.375 | 0.544 | 0.295 |

Hypertension Absent \& Present Group
Diabetes Absent \& Present Group



